

# Using Rasch Measures For Rasch Model Fit Analysis

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In Rasch fit analysis,  $Z_{ni}$  is used to measure the fit of a single person-item response, while mean-square (MS) statistics analyze the fit of response sets, and ZSTD tests the significance of a particular MS value.

Most analysts find the Rasch model person measures and item calibrations easier to understand and communicate than the  $Z_{ni}$ , MS, and ZSTD statistics. For instance, only through the necessary calculations do we know how much logit-misfit is involved for a given  $Z_{ni}$  or MS value. Furthermore,  $Z_{ni}$ , MS, and ZSTD are nonlinear functions of Rasch model values (e.g.,  $B_n - D_i$ ).

This paper introduces a Rasch model fit statistic that enables the analyst to interpret fit of a response on the same scale as person measures and item calibrations. Essentially, this is accomplished by explicitly incorporating the logistic Rasch model in the fit statistics.

## RESPONSE-FIT INDEX FOR DICHOTOMOUS CHOICES

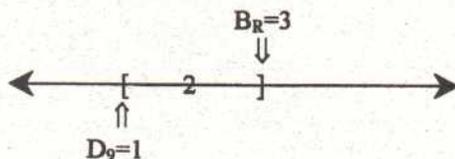
Let  $K_{ni}$  denote the logit-fit of person  $n$ 's response to item  $i$ , calculated by:  $K_{ni} = f_{ni}(B_n - D_i)$  [1]

where  $f_{ni}$  classifies the model-fit of a person-item response

$f_{ni} = 0$  for a response that fits the model  
( $X_{ni} = 1$  when  $B_n \geq D_i$ , or  $X_{ni} = 0$  when  $B_n < D_i$ )  
 $f_{ni} = -1$  for a response that misfits the model  
( $X_{ni} = 1$  when  $B_n < D_i$ , or  $X_{ni} = 0$  when  $B_n > D_i$ ).

**Example 1.** Richard with ability  $B_R = 3$  encounters "item 9" having difficulty  $D_9 = 1$ .

Map item and person on a number line:

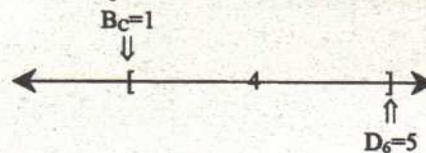


**Expected Response Rule:** Since  $B_n > D_i$ , then  $\{X_{ni} = 1\}$  is the expected response.

Two Possible Scenarios:

Response	Fit result	Interpretation
$\{X_{ni} = 1\}$	$K_n = 0(3-1) = 0$	Response fits measurement model.
$\{X_{ni} = 0\}$	$K_n = -1(3-1) = -2$	Richard responded 2 logits below expectation.

**Example 2.** Cindy with ability  $B_C = 1$  encounters "item 6" having difficulty  $D_6 = 5$ .

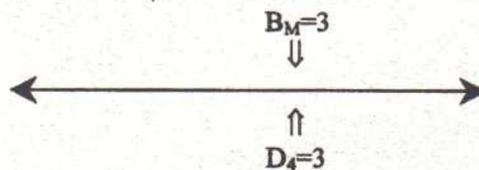


**Expected Response Rule:** Since  $B_n < D_i$ , then  $\{X_{ni} = 0\}$  is the expected response.

Two Possible Scenarios:

Response	Fit result	Interpretation
$\{X_{ni} = 1\}$	$K_{ni} = -1(1-5) = 4$	Cindy responded 4 logits above expectation.
$\{X_{ni} = 0\}$	$K_{ni} = 0(1-5) = 0$	Response fits measurement model.

**Example 3.** Mary with ability  $B_M = 3$  encounters "item 4" having difficulty  $D_4 = 3$ .



**Expected Response Rule:** Since  $B_n = D_i$ , then  $\{X_{ni} = 0\}$  and  $\{X_{ni} = 1\}$  have equal probability ( $P_{ni1} = .50$ , therefore  $P_{ni0} = .50$ ). So by definition, *neither response misfits the model.*

PSYCHOLOGICAL MEASUREMENT

Two Possible Scenarios:

Response	Fit result	Interpretation
$\{X_{ni}=1\}$	$K_{ni} = 0(3-3) = 0$	Response fits measurement model.
$\{X_{ni}=0\}$	$K_{ni} = 0(3-3) = 0$	Response fits measurement model.

RESPONSE-FIT INDEX FOR POLYTOMOUS CHOICES

Since all Rasch models reduce to the dichotomous-response model, Equation 1 can be extended to analyze the fit of a rating-scale response. For an item with  $m$  response categories, there are  $m-1$  adjacent-category steps, where each step  $j$  is denoted by the parameter  $F_j$ . A person's rating scale response to that item indicates a certain number of "advanced" steps, and a certain number of "unadvanced" steps. Each "advanced" versus "unadvanced" step response is a dichotomy, and therefore, there are  $j$  dichotomous responses within a single rating scale response.

The fit calculation of a single rating scale response involves calculating  $f_{nij}(B_n - D_i - F_j)$  for each of the steps, and letting  $K_{ni}$  equal the calculation that differs the most from zero. The  $K_{ni}$  for a single rating scale response is therefore calculated by:

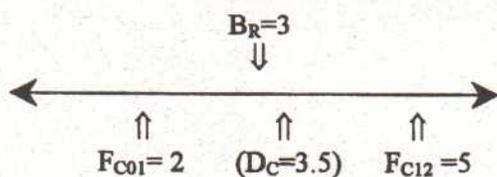
$$K_{ni} = |\max| [f_{nij} (B_n - D_i - F_j)] \quad [2]$$

where,  
 $|\max|$  maximum in absolute value  
 $f_{nij} = 0$  for a step-response that fits the model  
 $f_{nij} = -1$  for a step-response that misfits the model

In the case of dichotomous response choices, there is only one threshold  $j$ , in which case equation [2] reduces to equation [1].

Here is an example of an item with a three category ( $m=3$ ) rating scale, where  $X_{ni} = \{0,1,2\}$ , rendering  $m-1=2$  steps. Let  $F_{01}$  denote the parameter for the step to category 1 from 0, and  $F_{12}$  for the step to category 2 from 1.

*Example 4.* Bob with ability  $B_B=3$  encounters "item C" having difficulty  $D_C=3.5$ , where  $F_{01}=1.5$  and  $F_{12} = +1.5$  relative to  $D_C$ .



Expected Response Rule: Since  $B_n > F_{i01}$  and  $B_n < F_{i12}$ ,  $\{X_{ni}=1\}$  is the expected response.

Three Possible Scenarios:

Response	Fit result	Interpretation
$\{X_{ni}=2\}$	$K_{ni} =  \max  [0(3-2), -1(3-5)] = 2$	Bob responded 2 logits above expectation.
$\{X_{ni}=1\}$	$K_{ni} =  \max  [0(3-2), 0(3-5)] = 0$	Response fits measurement model.
$\{X_{ni}=0\}$	$K_{ni} =  \max  [-1(3-2), 0(3-5)] = -1$	Bob responded 1 logit below expectation.

FIT ANALYSIS OF RESPONSE SETS

Analyzing response sets is straightforward. The average of the absolute value of  $|K_{ni}|$  values can be taken across all responses of interest:

$$|\bar{K}_{ni}| = \frac{\sum |K_{ni}|}{N_{\{X_{ni}=x\}}} \quad [3]$$

to obtain the "average logit noise," where  $N_{\{X_{ni}=x\}}$  denotes the total number of responses. Person  $|\bar{K}_{ni}|$  is obtained by applying Equation 3 for all person responses; item  $|\bar{K}_{ni}|$  is calculated for all item responses.

It is also informative to take the average of certain response subsets. Examples include (1) the subset of "negative"  $K_{ni}$  values, and (2) the subset of "positive"  $K_{ni}$  values. Subset (1) indicates the magnitude of surprising "low" responses (e.g., occurring from sleeping, carelessness, etc.), and subset (2) indicates the magnitude of surprising "high" responses (e.g., lucky-guessing).

The accuracy of  $K_{ni}$  depends on parameter values estimated from the data, *but we know we estimate parameters from noisy data in the first place* ( $Z_{ni}$ , MS, ZSTD, and all parameter-dependent fit methods suffer this uncertainty). When data noise is high, we cannot trust the accuracy of parameter estimates, and therefore can no longer trust the accuracy of  $K_{ni}$  and other parameter-dependent fit statistics. In cases where data is too noisy for the parameter-dependent fit statistics to be useful, an alternative is an estimate of Guttman fit:

$$G = \frac{N_{|K_{ni}|>0}}{N_{\{X_{ni}=x\}}} \quad [4]$$

which is the proportion of unexpected responses across the relevant response set.  $G$  is linearized by the transformation  $\log(G/(1-G))$ .

It is also informative to change the numerator of Equation [4] to calculate the proportion of surprising "low" responses ( $N_{K<0}$ ) and "high" responses ( $N_{K>0}$ ).

$G$  interprets  $K_{ni}$  values as ordinal (possible values: either  $K_{ni}=0$  or  $|K_{ni}|>0$ ), which renders it more robust than  $|\bar{K}_{ni}|$  (and  $Z_{ni}$ , MS, ZSTD) to inaccurate parameter estimations. Hence,  $G$  can be considered a parameter-free fit statistic.

